

*Review Article*

# Enhancing Green Economy with Artificial Intelligence: Role of Energy Use and FDI in the United States

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## Abstract

The escalating challenge of climate change necessitates an urgent exploration of factors influencing carbon emissions. This study contributes to the discourse by examining the interplay of technological, economic, and demographic factors on environmental sustainability. This study investigates the impact of artificial intelligence (AI) innovation, economic growth, foreign direct investment (FDI), energy consumption, and urbanization on CO<sub>2</sub> emissions in the United States from 1990 to 2022. Employing the ARDL framework integrated with the STIRPAT model, the findings reveal a dual narrative: while AI innovation mitigates environmental stress, economic growth, energy use, FDI, and urbanization exacerbate environmental degradation. Unit root tests (ADF, PP, and DF-GLS) confirm mixed integration levels among variables, and the ARDL bounds test establishes long-term co-integration. The analysis highlights that AI innovation positively correlates with CO<sub>2</sub> reduction when environmental safeguards are in place, whereas GDP growth, energy consumption, FDI, and urbanization intensify CO<sub>2</sub> emissions. Robustness checks using FMOLS, DOLS, and CCR validate the ARDL findings. Additionally, Pairwise Granger causality tests reveal significant one-way causal links between CO<sub>2</sub> emissions and economic growth, AI innovation, energy use, FDI, and urbanization. These relationships emphasize the critical role of AI-driven technological advancements, sustainable investments, and green energy in fostering ecological sustainability. The study suggests policy measures such as encouraging green FDI, advancing AI technologies, adopting sustainable energy practices, and implementing eco-friendly urban development to promote sustainable growth in the USA.

**Keywords:** AI Innovation, CO<sub>2</sub> Emission, Energy Use, FDI, STIRPAT

## Introduction

The environment is crucial for supporting life on Earth; nevertheless, the increasing global pollution and waste have emerged as a significant problem in recent years [1,2]. The majority of climate-altering and planet-warming substances emitted into the atmosphere are carbon dioxide (CO<sub>2</sub>) [3]. The United States (USA) distinguishes itself among these nations through its evolution as a dominant force in consumption as well as production [4]. In 2020, the USA emitted 5,416 metric tons (MT) of CO<sub>2</sub>, or almost 16% of global emissions [5]. However, in 2021, the United States reinforced its dedication to the Paris Agreement, adopting an ambitious Nationally Determined Contribution to reach a 50–52% drop in net greenhouse gas (GHG) emissions by 2030 [6,7]. The United States consumes the largest quantity of fossil fuel annually, totaling 913.3 million tons of oil, which is 50% greater than China's consumption, the second highest globally [8,9]. Alleviating the adverse effects of global climate change has emerged as a worldwide priority, with a crucial component of this effort being the reduction of CO<sub>2</sub> emissions [10,11]. In accordance with the UN 2030 universal sustainable development objective, the advancement of renewable energy is prioritized above everything else. Consequently, the USA has considerable accountability for the climate crisis and global warming, being one of the foremost emitters of GHGs; thus, assessing its environmental sustainability is a critical issue. In this context, our research question is

to examine the response of GDP, FDI, energy consumption, urbanization, and AI innovation on CO<sub>2</sub> emissions in the context of United States.

Several scholars have thoroughly investigated the ecological conditions in the United States, mostly utilizing pollution metrics such as CO<sub>2</sub> emissions [12,13]. By encouraging domestic investment, facilitating technological transfers in the receiving nation, and boosting the development of human capital, foreign direct investment (FDI) enhances economic growth and is therefore essential to economic development [14]. Another viewpoint holds that FDI improves host countries' surroundings by bringing cutting-edge environmental technologies and sustainable practices [15,16]. The environment field is not an exception to the dramatic transformations we have lately experienced in numerous industries as artificial intelligence (AI) has grown more integrated. The United States leads in AI development and will be essential in developing suitable guardrails and regulatory frameworks that promote the responsible use of the technology [17]. The development of AI technologies can help address several global green growth concerns. Furthermore, applying AI can reduce environmental emissions [18]. Businesses are facing pressure to curtail GHG emissions and promote environmentally friendly practices as the global community battles climate change.

In 2020, the United States was the second-largest contributor to world pollution, producing 4.7 billion metric tons of CO<sub>2</sub>. Multiple studies has demonstrated the mixed consequences of GDP on CO<sub>2</sub> emission in different regions [19,20,21]. After acknowledging this fact, politicians and governments must create plans for resource sustainability and seek solutions that strike a balance between ecological sustainability and GDP [22]. Increased use of clean energy sources is thought to be able to lessen the negative economic effects of climate change [23,24]. Between 2005 and 2022, emissions from the energy sector decreased by 35.9%, during which the sector represented almost one quarter of total US emissions. Furthermore, the United States' share of GDP in 2015 that went toward renewable energy was 0.2%, far less than that of other developing nations like South Africa (1.4%), China (0.9%), India (0.5%), and Brazil (0.4%) [25]. With a 15.8% share of primary energy consumption and 13.8% share of CO<sub>2</sub> emissions in 2020, the United States is also the world's top primary energy consumer [26].

This study enhances existing research in the following manners: Firstly this paper critically analyzes the correlation among AI innovation, foreign direct investment, and CO<sub>2</sub> emissions in the United States. Secondly, unlike previous studies that solely examined developing nations, this research focuses on one of the world's most industrialized countries with a robust financial system, with the aim of investigating and clarifying the relationship between financial systems and environmental sustainability. Moreover, it is the inaugural investigation of the dynamic interconnections among AI innovation, energy consumption, and foreign direct investment (FDI). Specifically the United States provides an ideal context for investigating the correlation between economic activity and environmental sustainability, given its significant ecological deficit and developed economy. This study reveals that while AI innovation reduces CO<sub>2</sub> emissions, factors such as GDP, FDI, urbanization, and energy consumption exacerbate environmental degradation. This unique contribution is vital for stakeholders as it offers tangible insights into how industrialized countries, like the USA, may harmonize economic growth with environmental sustainability. Additionally, the research corroborates the STIRPAT framework along with sophisticated methods like ARDL in the United States, employing a contemporary dataset from 1990 to 2022. Furthermore, it employs a comprehensive approach and innovative econometric models to offer significant insights that aid the United States in its endeavors to attain SDG-7 and SDG-13, especially in the quest for carbon neutrality. This research offers governments, corporations, and environmental advocates scientifically substantiated ideas for reconciling economic and environmental objectives, chiefly via financial, technological, and industrial modifications.

The subsequent portions of this work are organized as follows: Section II examines the pertinent literature, Section III delineates the methodology and data, Section IV showcases the results and analysis, and Section V explores the policy implications and conclusions.

## **Literature Review**

It is clear from a thorough analysis of earlier research that not many studies have looked into the connection between GDP, energy use, FDI, CO<sub>2</sub> emissions, and AI innovation. The bulk of publications concentrated on the effects of trade openness, urbanization, and green energy usage on environmental quality, even though several investigations looked at the ARDL model. The USA hasn't seen much in-depth research on ecological degradation, a relatively young discipline. Nevertheless, the research drew upon a few previous studies to guide the selection of variables and methods. This section will address a select few of these questions.

### ***GDP and CO<sub>2</sub> emission***

Numerous researches have been conducted to ascertain the correlation between environmental and economic activity. For example, Manu and Sulaiman [27] examined the impact of energy consumption and economic growth on Malaysia's CO<sub>2</sub> emissions utilizing the OLS method for the years 1965 to 2015. Their research demonstrated that CO<sub>2</sub> emissions diminish as revenue escalates. The favorable relationship between economic expansion and CO<sub>2</sub> emission is found by Saudi et al.[28], and Zubair et al.[29]. Conversely, Sarkar et al. [30] analyzed time-series data from 1980 to 2016 to investigate the relationship among Malaysia's energy use, CO<sub>2</sub> emissions, and economic expansion. The empirical evidence indicated that usage of energy and economic expansion substantially elevated CO<sub>2</sub> emissions in Malaysia from 1980 to 2016. In order to examine the relationship between GDP and CO<sub>2</sub> emissions, Chen et al. [31] used China's annual data from 1990 to 2020 and the QARDL technique. They discovered that China's GDP has a positive effect on CO<sub>2</sub> emissions. Etokakpan et al. [32] analyzed the relationship between capital formation, globalization, CO<sub>2</sub> emissions, and GDP in Malaysia using a dataset from 1980 to 2014 within a multivariate framework. The authors employed an innovative combined co-integration test to ascertain the magnitude of the long-run equilibrium relationship. The empirical findings indicated that GDP adversely affected environmental quality. Multiple researcher such as Wada et al.[33] within Brazil, Rjoub et al. [34] in Turkey and He at al.[35] within Mexico found the same conclusions. On the other hand, Muhammad et al. [36] employed FMOLS and two-stage least squares regression methods to analyze the correlation between GDP growth and CO<sub>2</sub> emissions. The findings corroborated the U-shaped link indicated by the EKC theory in high- and upper-middle-income nations.

### ***AI innovation and CO<sub>2</sub> emission***

The loss of biodiversity and global warming are complex concerns that require advanced and inventive solutions [37]. Environmental professionals anticipate several benefits from AI tools [38]. Policymakers can use AI innovation to develop scientifically supported plans and strategies for green ecosystems [39]. Existing research indicates that investigations on carbon reduction related to AI are in their infancy. A pertinent study by Liu et al. [40] analyzes industrial robot data from 16 sectors in China from 2006 to 2016 to investigate the correlation between AI and energy intensity. The implementation of AI technology in the industrial sector diminishes energy intensity by enhancing industrial output while decreasing energy consumption and environmental degradation. Chen et al. [41] investigate the impact of AI on carbon emissions utilizing panel data from 270 Chinese cities spanning 2011 to 2017. Their empirical findings indicate that AI decreases carbon emissions by optimizing production processes, boosting communication facilities, and advancing green technological innovation. Negi [42] examines the investment trends in artificial intelligence originating from the three leading nations: China, India, and the United States. The paper delineates the measures the government has implemented to integrate artificial intelligence into its existing ecosystem. Green AI can enhance productivity and mitigate its adverse environmental impacts [43].

### ***Energy Use and CO<sub>2</sub> emission***

The primary contributor to climate change is the combustion of fossil fuels, which releases significant quantities of greenhouse gases into the environment [44]. But long-term cost reductions from using

alternative or green energy sources will raise people's overall level of living [45]. Raihan et al. [46] examine the relationship between energy use and CO<sub>2</sub> emissions in Malaysia from 1990 to 2019. They employed the ARDL and DOLS methodologies, which revealed a positive and significant energy use coefficient in relation to CO<sub>2</sub> emissions, indicating a 0.91% rise in CO<sub>2</sub> emissions for every 1% increase in energy usage. Adebayo and Kalmaz [47] employed ARDL, FMOLS, and DOLS estimators to reveal a substantial positive correlation between power consumption and CO<sub>2</sub> emissions in Egypt, utilizing data from 1971 to 2014. Moreover, Adebayo et al. [48] identified a positive correlation between CO<sub>2</sub> emissions and energy consumption by employing the ARDL model for MINT nations, covering the period from 1980 to 2018. Conversely, Raihan and Tuspekova [49] examined the correlation between green energy utilization and CO<sub>2</sub> emissions in Peru from 1990 to 2018. Using the DOLS and ARDL methodologies, they found a negative correlation and statistical significance in the utilization of clean energy, suggesting that a 1% increase in green power use results in a 0.52% reduction in CO<sub>2</sub> emissions over the long term. Multiple researcher such as Abbasi and Adedoyin [50] in China, Balsalobre-Lorente et al.[51] across BRICS countries, Saqib [52] in MENA region corroborated the destructive effect of energy usage on environment quality. On the other hand, Younis et al.[53], Salari et al.[54] concluded that energy consumption boost the environment quality. In the end, energy-efficiency regulations can enhance energy conservation, while increased investment in energy production and the promotion of energy savings will diminish carbon emissions [55].

#### ***FDI and CO<sub>2</sub> emission***

Foreign Direct Investment (FDI) can facilitate the transfer of cleaner technologies and sustainable practices, leading to enduring reductions in emissions [29,56]. Jafri et al. [57] investigate the asymmetric impact of FDI on CO<sub>2</sub> emissions utilizing the NARDL methodology for China from 1981 to 2019. Their findings indicate that a positive shift in FDI has a comparatively greater impact on CO<sub>2</sub> emissions. Moreover, several studies also found the same conclusion in different region [58,59,60,61]. On the other hand, Lin et al. [62] examine the impact of FDI on emission reduction in China from 2004 to 2015, employing geographic Durbin economic models with two-way fixed effects. The findings indicate that FDI facilitates a decrease in emissions nationwide. Furthermore, Eskeland and Harrison [63] contended that FDI typically accompanies energy-efficient technologies and may positively impact the natural world. Similar outcomes were also observed by Wang et al. [64] in China, Pata and Samour [65] within France, and Abbas et al.[66] in developing countries. Conversely, Haug and Ucal [67] shown that spikes in FDI had no statistically significant long-term effects on per capita CO<sub>2</sub> emissions.

#### ***Urbanization and CO<sub>2</sub> emission***

Economic growth and industrialization-driven urbanization are contributing to the rising utilization of fuels that generate greenhouse gas emissions [68].The concluding section of the research examines previous studies on the empirical relationship between urbanization and CO<sub>2</sub> emission. For instance, Mahmood et al. [69] examine the impact of urbanization on per capita CO<sub>2</sub> emissions in Saudi Arabia, analyzing data from the years 1968 to 2014. The findings indicate that urbanization hinders the environment due to its elastic impact on emissions. Parshall et al. [70] examined the substantial impact of urbanization on the circular economy and environmental health in the USA. Raihan et al. [71] explore the impact of urbanization on the load capacity factor in Mexico from 1971 to 2018. This study utilizes the ARDL approach and demonstrates that urbanization decreases Mexico's LCF, thereby degrading the ecology. Sufyanullah et al. [72] examined the impact of urbanization on CO<sub>2</sub> emissions in Pakistan. They employed the ARDL model and noted that CO<sub>2</sub> emissions rise with increased urbanization. Additional studies [73,74,75,76,77] suggest that urbanization increases CO<sub>2</sub> emissions in the atmosphere. In comparison, Xu et al. [78] evaluated the impact of urbanization on the ecosystem in Brazil from 1970 to 2017. Unexpectedly, the results of the ARDL methodology indicated that urbanization does not influence the surroundings in Brazil. Moreover, Haseeb et al. [79] revealed analogous findings utilizing FMOLS

from 1995 to 2014, suggesting that URB had no significant impact on environmental quality in the BRICS nations. Nonetheless, as highlighted by Martinez et al. [80] urbanization may contribute to addressing climate change due to the heightened knowledge among these groups. Moreover, Diputra and Baek [81] determined that urbanization exerted no substantial impact on emissions in Indonesia.

**Literature Gap**

The current status of research indicates that studies on the USA are scarce. The empirical literature lacks an evaluation of the geographical implications of macroeconomic variables on CO<sub>2</sub> emissions in the USA. The literature on the STIRPAT framework and ARDL methodology is deficient in the USA and similarly limited in other global locations. This study tries to fill in that gap by looking at modern variables using advanced methods from a U.S. point of view. The goal is to find the main and secondary effects of GDP, AI innovation, energy consumption, urbanization, and FDI on CO<sub>2</sub> emissions in this area. By examining these procedures, the USA may determine whether leveraging technological innovation, financial integration, and commercial expansion can enhance its ecosystem quality and align it with global trends toward greater environmental sustainability.

**Methodology**

**Data and Variables**

This study seeks to observe the impact of GDP, FDI, AI innovation, energy use, and urbanization on CO<sub>2</sub> emissions in the USA. The study utilizes CO<sub>2</sub> emissions as an indicator of ecological health, employing data from the World Development Indicators (WDI) database as the dependent variable. Statistics on AI innovation are sourced from Our World in Data, while data for GDP, energy consumption, foreign direct investment (FDI), and urbanization are derived from the revised WDI. These variables encompass annual data from 1995 to 2022. Table 1 presents a complete enumeration of each variable along with their respective details and a sign chosen for this research.

Table 1. Variables description

Variables	Description	Logarithmic Form	Unit of Measurement	Source
CO <sub>2</sub>	CO <sub>2</sub> Emission	LCO <sub>2</sub>	CO <sub>2</sub> Emission (kt)	WDI
GDP	Gross Domestic Product	LGDP	GDP per capita (current US\$)	WDI
AI	AI Innovation	LPAI	Estimated Investment in AI (US\$)	Our World in Data
ENU	Energy use	LENU	Energy use (kg of oil equivalent per capita)	WDI
FDI	Foreign Direct Investment	LFDI	Net Inflows (Current US\$)	WDI
URB	Population	LPOP	Population, total	WDI

**Theoretical Framework**

The STIRPAT paradigm holds substantial importance for environmental research. This method is a versatile analytical tool that facilitates the understanding of intricate relationships between human societies and the environment, irrespective of the subject matter, including GHG emissions, air pollution,

deforestation, or biodiversity loss [82]. By using STIRPAT, this study allows for an extensive evaluation of numerous variables that affect CO<sub>2</sub> emissions in the context of USA. The IPAT model was proposed by Ehrlich and Holdren [83] and is phrased as follows:

$$I = \int PAT \tag{1}$$

Nonetheless, this model proved challenging to evaluate hypothetically [84]. To address these constraints, Dietz and Rosa [85] expanded the IPAT model, resulting in the STIRPAT model. This research can effectively assess the interplay between human activities and environmental outcomes by representing CO<sub>2</sub> emissions as the 'I' or impact variable, urbanization as the 'P', and other variables like AI innovation, GDP, energy consumption, as the 'T' or technological variables. Equation (2) shows the functional form of the STIRPAT framework:

$$I_i = C \cdot P_i^\alpha \cdot A_i^\beta \cdot T_i^\gamma \cdot \varepsilon_i \tag{2}$$

Based on a comprehensive review of relevant literature, the empirical model employed in this work yielded the subsequent approximations.

$$\text{Environmental Impact} = f(\text{Population, Affluence, Technology}) \tag{3}$$

To assess the effects on the environment, this study uses CO<sub>2</sub> emissions as a proxy indicator. This is the expression that may be used to derive Equation (4):

$$CO_{2it} = \delta_0 + \delta_1 GDP_{it} + \delta_2 AI_{it} + \delta_3 ENU_{it} + \delta_4 FDI_{it} + \delta_5 URBA_{it} + \varepsilon_{it} \tag{4}$$

Here, GDP stands for gross domestic product; AI innovation was represented by AI, energy use through ENU, foreign direct investment via FDI and urbanization by URBA. Equation (5) makes use of the logarithmic transformation of variables to ensure that the information has a normal distribution. By employing the logarithmic structure, the info is standardized which makes it more consistent with the assumptions that underlie numerous statistical methods.

$$LCO_{2it} = \delta_0 + \delta_1 LGDP_{it} + \delta_2 LAI_{it} + \delta_3 LENU_{it} + \delta_4 LFDI_{it} + \delta_5 LURBA_{it} + \varepsilon_{it} \tag{5}$$

**Empirical Methods**

This study utilized the ARDL approach for data analysis to explore the correlation between CO<sub>2</sub> emissions and several independent variables in the USA. To ensure robustness, the Fully Modified Ordinary Least Squares (FMOLS), Dynamic Ordinary Least Squares (DOLS), and Canonical Cointegration Regression (CCR) methods were additionally employed. Initially, the author performed unit root tests (ADF, P-P, and DF-GLS) to verify stationarity. The properties of the time series data led to the employment of the ARDL limits test. Subsequently, both short-term and long-term ARDL estimates were derived, followed by the Pairwise Granger causality examination. Ultimately, multiple diagnostic assessments were conducted, allowing us to find the most precise and effective econometric model following a comprehensive evaluation procedure.

**Unit root test**

Performing a unit root testing is needed to avert erroneous regression analysis. This test can determine the degree of integration [86]. The Dickey-Fuller Generalized Least Squares, [87] Phillips-Perron [88], and Augmented Dickey-Fuller [89] unit root tests were used to see if the data set was stationary. In

comparison to the Dickey-Fuller (DF) method, the Augmented Dickey-Fuller (ADF) technique is more resilient and suitable for more complex procedures [90]. The DF-GLS test shows superior overall performance regarding small sample size and power, surpassing the conventional Dickey-Fuller test [91]. Before employing the ARDL bound test estimator for cointegration analysis, it is essential to conduct the ADF and PP unit root tests, as the estimator is applicable only when the variables are stationary at the level or first difference [92].

### **ARDL technique**

The ARDL technique, which includes the features of both distributed lag and autoregressive models, was proposed by Pesaran et al. [93]. It is a comprehensive dynamic regression model that offers various advantages over traditional cointegration methods. Firstly, the model allows for the integration of variables to multiple orders, including order one, order zero, and even fractional integration, with the exception of 1(2). Moreover, unlike previous cointegration approaches that necessitated prior identification of a series' integration characteristic, this approach does not mandate any preliminary validation [94]. Due to its efficiency, we can employ this methodology for data analysis in scenarios with limited and small sample sizes [95,96]. Equation (6) can be used to represent the ARDL bound test:

$$\begin{aligned} \Delta LCO_{2t} = & \beta_0 + \beta_1 LCO_{2t-1} + \beta_2 LGDP_{t-1} + \beta_3 LAI_{t-1} + \beta_4 LENU_{t-1} + \beta_5 LFDI_{t-1} \\ & + \beta_6 LURBA_{t-1} + \sum_{i=1}^m \alpha_1 \Delta LCO_{2t-i} + \sum_{i=1}^m \alpha_2 \Delta LGDP_{t-i} + \sum_{i=1}^m \alpha_3 \Delta LAI_{t-i} \\ & + \sum_{i=1}^m \alpha_4 \Delta LENU_{t-i} + \sum_{i=1}^m \alpha_5 \Delta LFDI_{t-i} + \sum_{i=1}^m \alpha_6 \Delta LURBA_{t-i} + \varepsilon_t \end{aligned} \tag{6}$$

Where, m is the optimum lag length. Pesaran et al. [93] suggest using critical values for both upper and lower bounds to compare F-statistics. When the F-statistic surpasses the upper critical value, we reject the null hypothesis (H0), indicating a persistent relationship. If the F-statistic stays below the crucial value, we retain the null hypothesis (H0). The long-run coefficient estimate is derived from equation (7), which also validates the cointegration of the parameters. It uses an approximation of the Error Correction Term (ECT) to figure out short-term dynamic parameters based on long-term estimates [97]. The ECT is built into the ARDL structure. The equation (7) outlines the ARDL long-run equation presented below.

$$\begin{aligned} \Delta LCO_{2t} = & \beta_0 + \beta_1 LCO_{2t-1} + \beta_2 LGDP_{t-1} + \beta_3 LAI_{t-1} + \beta_4 LENU_{t-1} + \beta_5 LFDI_{t-1} \\ & + \beta_6 LURBA_{t-1} + \sum_{i=1}^m \alpha_1 \Delta LCO_{2t-i} + \sum_{i=1}^m \alpha_2 \Delta LGDP_{t-i} + \sum_{i=1}^m \alpha_3 \Delta LAI_{t-i} \\ & + \sum_{i=1}^m \alpha_4 \Delta LENU_{t-i} + \sum_{i=1}^m \alpha_5 \Delta LFDI_{t-i} + \sum_{i=1}^m \alpha_6 \Delta LURBA_{t-i} \\ & + \Omega ECT_{t-1} + \varepsilon_t \end{aligned} \tag{7}$$

### **Robustness Check**

To evaluate the precision of ARDL outcomes, we utilized the FMOLS, DOLS, and CCR techniques. The FMOLS approach is utilized to analyze a singular cointegrating relationship involving a combination of integrated orders of I(1) variables. The primary objective of this method is variable transformation. Phillips and Hansen [98] say that the FMOLS method fixes the problems with traditional cointegration methods that make it hard to draw conclusions. This makes the estimated t-statistics for long-term estimates more reliable. The DOLS technique may assist in integrating individual variables within a cointegrated framework when faced with a mixed order of integration. The dependent variable is calculated utilizing levels, leads, and lags as explanatory variables [99]. Nevertheless, as emphasized by

Pesaran [100], the principal advantage of the DOLS prediction is its allowance for the varying order integration of distinct components within the cointegrated framework. Additionally, Park [101] proposed the CCR technique to examine cointegrating vectors within a model characterized by an integrated process of order I(1). The model's characteristics exhibit a significant similarity to FMOLS.

**Pairwise granger causality**

The study employed the paired Granger-causality test, devised by Granger [102], to ascertain the existence of a causal relationship among the factors. We can use F tests to assess Granger causality between variables X and Y, and the OLS test for coefficient estimation. The symbols  $X_t$  and  $Y_t$  denote the values of the variables at time t, illustrating the time series for this variable pair. A bivariate autoregressive model may exhibit the variables  $X_t$  and  $Y_t$ .

$$X_t = \beta_1 + \sum_{i=1}^n \alpha_i Y_{t-i} + \sum_{i=1}^n \mu_i X_{t-i} + e_t \tag{8}$$

$$Y_t = \beta_2 + \sum_{i=1}^n \Omega_i Y_{t-i} + \sum_{i=1}^n \infty_i X_{t-i} + u_t \tag{9}$$

Here, the information criterion determines the "n" number of lags. The parameters used for the assessment were  $\beta_1$ ,  $\beta_2$ ,  $\alpha_i$ ,  $\Omega_i$ ,  $\mu_i$ , and  $\infty_i$ .

**Diagnostic test**

The errors in Equation (7) must not exhibit serial correlation. This study employed various diagnostic techniques to verify the normality, serial correlation, and heteroscedasticity of the data. Three tests are needed in time series analysis to make sure that model assumptions are correct and that results are stable: the Lagrange Multiplier (LM) test, the Jarque-Bera test [103], and the Breusch-Pagan-Godfrey test [104]. The Jarque-Bera test assesses the normality of residuals, a crucial step since many econometric models require normally distributed errors for precise inference. The Lagrange multiplier test examines residuals for serial correlation to verify that errors do not correlate with time, thereby preventing biased and misleading estimates. The Breusch-Pagan-Godfrey test can yield inaccurate estimates and standard errors due to heteroscedasticity, or the non-constant variance of residuals. The model's stability was assessed by CUSUMSQ studies [105].

**Results and Discussion**

Table 2 lays out the descriptive statistics for the considered variables, derived from 32 observations. The table provides the mean, standard deviation, minimum, and maximum values for six variables in the USA: LCO2, LGDP, LAI, LENU, LFDI, and LURBA. All examined variables demonstrate positive mean values, with LCO2 showing the highest mean. LFDI has a minimum value of 2.268, while LURBA has the highest number.

Table 2. Summary Statistics.

Variable	Obs	Mean	Std. Dev.	Min	Max
T	32	2005.5	9.381	1990	2021
LCO2	32	15.464	.08	15.279	15.569
LGDP	32	10.644	.319	10.081	11.159
LAI	32	7.506	1.036	6.321	9.724
LENU	32	4.77	.322	3.949	5.272
LFDI	32	2.625	.133	2.268	2.871
LURBA	32	19.5	.087	19.335	19.621



Moreover, all variables have relatively small standard deviations, indicating a tight clustering of data points around the mean with limited temporal variability. Table 3 presents the results of all three stationarity tests (ADF, DF-GLS, and P-P) for log-transformed data in both level I(0) and first-difference I(1) forms. In all three unit root assessments, it seems that only the urbanization variable is stationary at level I(0), whereas CO<sub>2</sub>, GDP, AI innovation, energy consumption, and FDI were non-stationary prior to examining their initial differences. This mixed sequence of integration prompts us to proceed with the assessment immediately, using the ARDL methodology.

Table 3. Results of unit root tes.t

Variables	ADF		P-P		DF-GLS		Decision
	I(0)	I(1)	I(0)	I(1)	I(0)	I(1)	
LCO <sub>2</sub>	-0.233	-3.941***	-0.231	-4.001***	-0.234	-3.991***	I(1)
LGDP	-0.872	-4.091***	-0.782	-4.891***	-0.809	-4.091***	I(1)
LAI	-0.704	-5.105***	-0.802	-5.323***	-0.756	-5.034***	I(1)
LENU	-0.172	-5.011***	-0.177	-5.071***	-0.819	-2.150***	I(1)
LFDI	-0.072	-4.108***	-0.065	-4.015***	-0.025	-4.342***	I(1)
LURBA	-5.012***	-7.011***	-5.801***	-7.605***	-5.831***	-7.050***	I(0)

Standard errors in parentheses  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

This research employed the ARDL bounds testing approach to ascertain the presence of co-integration among the variables. The F-statistic of 5.3421 is more than the critical value, indicating that the null hypothesis of no co-integration is rejected at the 1% significance level. Therefore, we can argue that the parameters of the model exhibit co-integrating relationships. This study cites urbanization, artificial intelligence innovation, foreign direct investment, gross domestic product, and energy consumption as the enduring driving forces. Furthermore, these factors necessitate the system's initial response to a typical stochastic disturbance. In summary, variations in these three parameters influence CO<sub>2</sub> emissions in the United States.

Table 4. Results of ARDL bound test.

Test Statistic	Value	Signif.	I(0)	I(1)
F-statistic	5.3421	10%	2.07	3
k	5	5%	2.43	3.27
		2.50%	2.81	3.84
		1%	3.10	4.20

Table 4 and table 5 adopts the dynamic ARDL model to demonstrate the short- and long-term effects of LGDP, LAI, LENU, LFDI, and LURBA on LCO<sub>2</sub> in the USA. In terms of LGDP, a 1% boost in LGDP will increase the LCO<sub>2</sub> by 0.028% in the long-term run and 0.012% in the short run. This suggests that economic expansion alone may notably contribute to environmental degradation in this setting, as GDP has a positive impact on the CO<sub>2</sub> emission level. A few studies have concluded that a boost in the GDP has a negative impact on the environment. This includes Ahmed et al. [106] in Japan; Raihan and Tuspekova [107] within Kazakhstan; Orhan et al. [108] on India; Ali et al. [109] in Malaysia; Shang et al.

[58] for ASEAN countries. However, Le and Ozturk [110], Zhan et al.[111], Tufail et al.[112] and He et al.[113] discovered the opposite outcome. They also concluded that economic pressures no longer adversely affect the natural world. Likewise, Awosusi et al. [114] demonstrated that there is no significant correlation between CO<sub>2</sub> emissions and GDP in MINT economies.

On the other hand, the coefficients for LAI indicate a positive correlation with LCO<sub>2</sub>, implying a 0.054% long-term fall and 0.076% short-term cut in LCO<sub>2</sub> for every 1% rise in PAI. Thus, private investment in artificial intelligence in the United States significantly contributes to environmental sustainability. AI enhances energy efficiency and promotes environmental sustainability by decreasing carbon emissions. Multiple researchers like Nishant et al. [37], Chen et al.[31] and Zhao et al.[115] support this result, stating that AI technologies improve the environmental conditions in different regions. Moreover, Wang et al. [116] discovered that invention patents exhibited no significant correlation with emissions decrease. Conversely, LCO<sub>2</sub> is negatively associated with LENU in both the long and short run, and this relationship is statistically significant. These findings suggest that energy consumption has an adverse impact on the USA ecosystem. Specifically, a 1% increase in LENU increases LCO<sub>2</sub> by 0.617% in the long run and by 0.321% in the short run. The utilization of energy results in increased carbon emissions, as fossil fuels, the predominant energy source, emit substantial CO<sub>2</sub> during combustion. This result is consistent with the research of Islam et al. [117] in Bangladesh, Kim [118] in OECD countries, Nurgazina et al. [119] in Malaysia, Akbota and Baek [120] within Kazakhstan, and Odugbesan and Adebayo [121] in Nigeria. On the other hand, the findings of Namahoro et al. [122], Bhat [123], and Sharif et al. [4] concluded that energy use can have a negative impact on the environment by increasing the pollution level.

Table 5. Results of ARDL short-run and Long-run.

Variable	Coefficient	Std. Error	t-Statistic	Prob.
Long-run Estimation				
LGDP	0.028	0.5432	0.0771	0.010
LAI	-0.054	0.0140	-2.3501	0.003
LENU	0.617	0.1177	1.4516	0.012
LFDI	0.018	0.0253	2.6532	0.014
LURBA	0.710	0.5406	2.0061	0.021
C	10.872	4.0321	3.0562	0.000
Short-run Estimation				
D(LCO2(-1))	0.517	0.1023	1.2054	0.041
D(LGDP)	0.012	0.3621	5.0452	0.000
D(LAI)	-0.076	0.0054	-3.1802	0.025
D(LENU)	0.321	0.1072	2.5638	0.024
D(LFDI)	0.031	0.1892	-4.6723	0.000
D(LURB)	0.641	1.4912	2.6732	0.022
CointEq(-1)*	-0.398	0.1040	-4.4572	0.003

Standard errors in parentheses  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Similarly, there is an unfavorable correlation between LFDI and LCO<sub>2</sub>, with each 1% increase in FDI increases the CO<sub>2</sub> emission by 0.018% in the long run and 0.031% in the short run. This result is significant at the 1% level. One possible reason is that FDI frequently results in elevated industrial activity, increased energy consumption, and intensified resource exploitation; hence, it amplifies environmental deterioration. Nie et al. [124], and Zhang et al. [125] corroborate this result. Conversely, Azam and Raza [126], Shah et al.[127], Paziienza [128] and Pradhan et al. [129] revealed that FDI can mitigate CO<sub>2</sub> emissions and improve biodiversity quality. Additionally, the positive and statistically significant URBA coefficients indicate that both long-term and short-term increases in LURBA negatively affect environmental quality. A 1% increase in URBA raises LCO<sub>2</sub> by 0.710% in the long run and by 0.641% in the short run. These findings suggest that the current urbanization structure in the United States is not conducive to reducing pollution. Several researchers have also observed a similar outcome, including Yuan et al. [130] in China, Ali et al. [131] in Pakistan, Anwar et al. [132] in Asian economies, Sikder et al.[133] in developing economies, and Raihan et al. [71]. However, studies conducted by Wang et al.[76], Acheampong [12], Zhu et al. [134], and Gasimli et al. [135] have demonstrated that urbanization enhances environmental sustainability by reducing carbon dioxide emissions.

The DOLS, FMOLS, and CCR techniques are supplementary methodologies utilized to evaluate the validity and reliability of the ARDL results. Table 6 outlines the robustness findings.

In the FMOLS model, the coefficients for LGDP are statistically significant at the 1% level and have positive values. A 1% increase in GDP causes the LCO<sub>2</sub> to rise by 0.245%. Additionally, a 1% increment in LAI leads to a 0.034% drop in CO<sub>2</sub> emission in the USA. Furthermore, a 1% boosts in LENU and LFDI and LURBA upsurges LCO<sub>2</sub> by 0.074%, 0.053% and 0.231%, respectively. It indicates that GDP, energy consumption, FDI and urbanization are not good for better for the ecosystem in USA. These findings corroborate the ARDL short and long-run estimation results, with LGDP, LAI, LENU and LFDI significant at the 1% level, while LURBA are significant at the 5% level. In the DOLS model, a 1% spike in LGDP, LENU, LFDI and LURBA results in an average rise of 0.316%, 0.023%, 0.039%, and 0.143% in LCO<sub>2</sub>, respectively. Similar to the ARDL findings, a 1% rise in LAI leads to a 0.010% reduction in LCO<sub>2</sub>, and the coefficient for LAI is significant at the 5% level.

In the CCR model, a 1% increase in LGDP, LENU, LFDI and LURBA leading to an average rise of 0.217%, 0.354%, 0.049%, and 0.205% in LCO<sub>2</sub>, respectively. However, a 1% increase in LAI causes an average 0.037% decrease in LCO<sub>2</sub>, confirming the ARDL results except for the LAI case. In this case all the factors are significant at 1% level, while LURBA is significant at 5% thresholds. These robustness checks confirm that the ARDL model's findings are reliable, as evidenced by the statistically significant values across FMOLS, DOLS, and CCR computations.

Table 6. Results of Robustness check.

Variables	FMOLS	DOLS	CCR
LGDP	0.245***	0.316***	0.217***
LAI	-0.034***	-0.010**	-0.037***
LENU	0.074***	0.023***	0.354***
LFDI	0.052***	0.039**	0.049***
LURB	0.231**	0.143**	0.205**
C	10.342***	10.052***	10.034***

Standard errors in parentheses  
 \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 7 presents the conclusions of causal relationships among several economic indices. An F-statistic of 3.38423 and a p-value of 0.0102 suggest that LLGDP does not cause Granger-cause LCO<sub>2</sub>. This indicates the rejection of the null hypothesis asserting no correlation between the variables at the 1% significance level. Also, p-values below the usual level of significance supported the finding that LAI, LENU, LFDI and LURBA all had a single-direction effect on LCO<sub>2</sub>. Consequently, under these conditions, we dismiss the null hypothesis, asserting the absence of causation. On the other hand, p-values higher than the usual significance level showed that there was no significant causal relationship from LCO<sub>2</sub> to LGDP, LAI, LENU, LFDI and LURBA. The null hypothesis, which posits the absence of causality in these interactions, is not effectively disproved.

Table 7. Results of Pairwise Granger Causality test

Null Hypothesis	Obs	F-Statistic	Prob.
LGDP ≠ LCO2	30	3.8423	0.0102
LCO2 ≠ LGDP		0.7345	0.1203
LAI ≠ LCO2	30	3.0201	0.0073
LCO2 ≠ LAI		0.0123	0.8341
LENU ≠ LCO2	30	4.6512	0.0143
LCO2 ≠ LENU		0.6712	0.7612
LFDI ≠ LCO2	30	4.8061	0.0041
LCO2 ≠ LFDI		0.7623	0.0871
LURB ≠ LCO2	30	4.7032	0.0191
LCO2 ≠ LURB		0.4121	0.5412

The results of the diagnostic evaluation are shown in Table 8. The findings indicated that the efficacy of all diagnostic techniques is negligible, and the null hypothesis remains indisputable. The Jarque-Bera test, yielding a p-value of 0.4321, suggests a normal distribution of the residuals. The Lagrange multiplier analysis indicates the absence of serial correlation in the residuals, yielding a p-value of 0.1021. Finally, the Breusch-Pagan-Godfrey test indicates that the residuals do not display heteroscedasticity, yielding a p-value of 0.1283.

Table 8. The results of diagnostic tests

Diagnostic tests	Coefficient	p-value	Decision
Jarque-Bera test	1.2034	0.4321	Residuals are normally distributed
Lagrange Multiplier test	1.0982	0.1021	No serial correlation exists
Breusch-Pagan-Godfrey test	0.0452	0.1283	No heteroscedasticity exists

Furthermore, we use the CUSUM and CUSUM-SQ statistics to seek structural stability in residuals over long and short periods. The CUSUM-SQ plot stays on the critical line, as indicated in the figure 1, indicating that the results are within the critical limits. This shows that the parameters are satisfactory and consistent at the 5% level of significance.

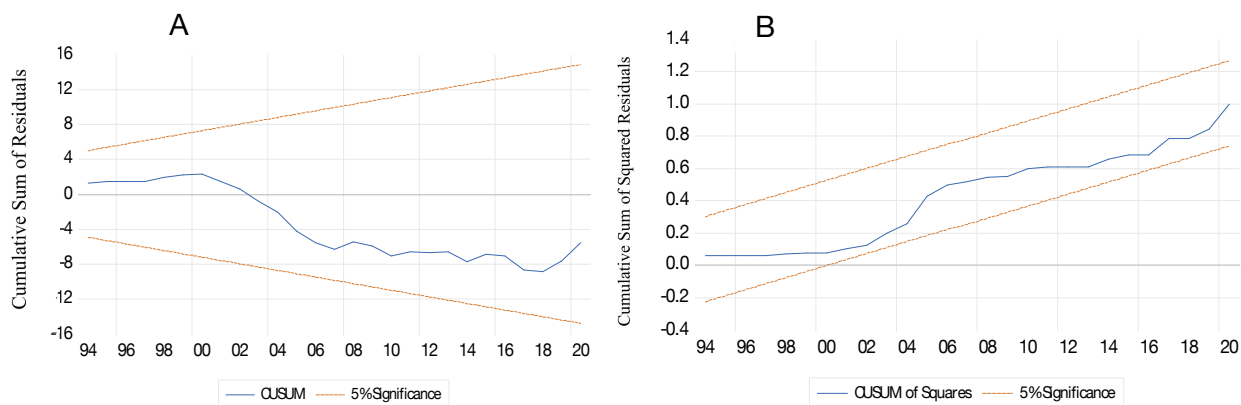


Figure 1. (A) CUSUM test for structural stability in residuals, (B) CUSUM of squares test for structural stability in residuals.

### Conclusion and Policy Implications

This paper thoroughly examines the impact of AI innovation, economic growth, foreign direct investment, energy consumption, and urbanization on CO<sub>2</sub> emissions in the USA from 1990 to 2022. Combining the ARDL framework with the STIRPAT structure, the authors found that while AI innovation reduces stress on the environment, economic growth, energy use, foreign direct investment, and urbanization make these problems worse. The results of the ADF, PP, and DF-GLS tests show that the variables have different levels of integration and no unit root problems. The ARDL boundaries test offers additional evidence of co-integration, signifying solid long-term interactions. The ARDL findings show a positive correlation between AI innovation and CO<sub>2</sub> emissions in the USA, suggesting that AI technologies improve environmental health as long as appropriate environmental safeguards are in place. In contrast, the adverse correlations among GDP, ENU, FDI, URBA, and CO<sub>2</sub> emissions indicate that these elements lead to detrimental environmental consequences. Significant technological advancements, including energy conservation, sustainable foreign direct investment, and urban planning, may stimulate new concepts and the implementation of environmentally friendly practices by enhancing competitiveness and facilitating access to advanced technologies. Robustness checks using FMOLS, DOLS, and CCR enhance the credibility of the ARDL results, thereby increasing their trustworthiness. Another thing is that Pairwise Granger causality tests show strong one-way links between LCO<sub>2</sub> and LGDP, LAI, LENU, LFDI, and LURBA. These linkages highlight the significant impact of economic transformations, private investments in AI, and advancements in green energy utilization on ecological sustainability dynamics in the USA. The research suggests various policy measures to promote sustainable economic growth in the United States, including the use of foreign direct investment, technological advancements, the implementation of green energy, and the development of sustainable urban infrastructure.

The findings emphasize the need for targeted policies to balance economic growth with environmental sustainability. Promoting AI-driven technologies and green energy initiatives can reduce CO<sub>2</sub> emissions while fostering innovation. Policies should prioritize sustainable foreign direct investment (FDI) by incentivizing eco-friendly projects and encouraging energy-efficient practices. Urban planning reforms must focus on developing smart, sustainable cities to mitigate the environmental impacts of urbanization. Additionally, integrating AI in energy management and industrial processes can enhance efficiency and reduce environmental stress. These measures collectively support sustainable economic growth while safeguarding ecological health in the USA.

### **Declaration**

**Ethics approval/declaration:** Not applicable.

**Consent to participate:** Not applicable.

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**Data availability:** Data will be available upon reasonable request from corresponding author.

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